

methods

Assessing the Validity of the Standardized Assessment of Reading Skills in Russian and Verifying the Relevance of Available Normative Data

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Abstract. Standardized tests with normative data have become the gold standard for assessing reading skills and for diagnosing a specific reading disorder (developmental dyslexia) in many languages. For Russian, there is one such reading assessment test called the Standardized Assessment of Reading Skills, developed by A.N. Kornev and first published in 1997. However, the most recent available normative data on this methodology were collected more than a decade ago, and researchers did not control for several important variables. Furthermore, no details have been published about the diagnostic validity of this test. We used the test to assess reading skills in 90 typically developing Russian primary school children in 2018. In this article, we present the results of testing typically developing children, including updated values of reading fluency and, for the first time, metrics of reading comprehension and weighted error scores. Additionally, we tested 50 children with clinically diagnosed developmental dyslexia and provide information about the sensitivity and specificity of the Standardized Assessment of Reading Skills.

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Introduction

Standardized tests are necessary for reliable and valid quantification of reading abilities, and for comparison between groups and individuals in order to diagnose reading disorders. Reading tests with uniform administration protocol and normative data from a large sample of individuals have become a methodological gold standard in last decades. Complex reading and spelling tests were developed for the German language (Moll & Landerl, 2010; Wimmer & Mayringer, 2014), standardized word-reading test (Brus & Voeten, 1973) and standardized pseudoword-reading test (van den Bos, Spelberg, Scheepstra, & DeVries, 1994) were developed for Dutch. For English, several variants of standardized reading assessments are available, including specialized reading test batteries (Torgesen, Wagner, & Rashotte, 1999; Wechsler, 1990; Woodcock, 1999) and reading subtests of more comprehensive cognitive assessment tools, such as Wide Range Achievement Test (Wilkinson & Robertson, 2006). A standardized reading test is also available for Russian.

The Standardized Assessment of Reading Skills (SARS) for Russian-speaking children was developed in 1982, first published in 1997 (Kornev, 1997), and republished in 2003 (Kornev, 2003). The second round of normative data collection was carried out in 2007–2008 (Kornev & Ishimova, 2010) to capture changes in the mean reading performance of children which happen over time. This test, according to the author, is intended for the “diagnosis of reading disorders and differential diagnosis of developmental dyslexia and non-specific reading disorders” (Kornev & Ishimova, 2010). Other existing reading assessment tools for Russian are parts of comprehensive neuropsychological assessment protocols (Akhutina et al., 2016; Semenovich, 2002), but their diagnostic application is limited due to the lack of published normative data.

The SARS has a number of obvious strengths. Firstly, during restandardization the number of children tested for normative data was substantially increased. First edition (Kornev, 1997) was published with normative data of 150 schoolchildren. The current edition (Kornev & Ishimova, 2010) was published with normative data of 700 schoolchildren. Secondly, this reading test provides the opportunity to measure not only reading fluency, but also reading comprehension using content questions with straightforward guidelines for scoring the responses as correct or incorrect. Thirdly, it includes texts of varying

difficulty (simple Text I and more difficult Text II), which allows to compare the performance of a reader as a function of increasing text difficulty. Kornev and Ishimova (2010) report that 10 and more points difference between reading performance of the two texts indicates greater severity of reading disorder. Finally, the SARS includes additional texts intended for retesting if necessary, with levels of difficulty comparable to the first two texts, and also with normative data of 400 Moscow schoolchildren tested in 2009–2010.

However, the SARS in its present form has limitations. Randomly chosen children from public schools took part in testing for normative data collection (Kornev & Ishimova, 2010), while it is preferable to include children without diagnosed neurological disorder or intellectual disabilities to determine appropriate cutoff scores for normal performance (Ivanova & Hollowell, 2013). In addition, despite the provided content questions for assessing reading comprehension and the well-developed criteria for scoring responses as correct or incorrect, the authors have not published normative data for reading comprehension. They suggested considering the level of performance with at least 7 out of 10 correct responses as “functionally sufficient”. At the same time, they provided information on the distribution of reading comprehension levels in the tested Moscow schoolchildren. From these data of the SARS’ authors, it can be calculated that 53% and 69% of second-graders correctly answered fewer than 7 questions for Text I and Text II, respectively (Kornev & Ishimova, 2010). Thus, the suggested level-based recommendation cannot serve as a guide for diagnosing reading comprehension disorder. Normative scores for reading comprehension are required, similarly to reading fluency.

Even more importantly for diagnostic purposes, in the last edition of the SARS (Kornev & Ishimova, 2010), direct normative data, e.g., the mean reading fluency and the standard deviation, were not provided. Instead, the authors suggested a transformation with a special formula resulting in coefficients of reading technic (see Method section). Those coefficients were not given for every possible number of correctly read words, but with gaps of 3 points. As a result, when the performance of a particular child is at the borderline (e.g., between “at risks” and “dyslexia”), there is no definite answer to which group the performance should be assigned to. Finally, Kornev (2003, p. 213) claimed that “the test showed sufficient validity in the clinical population”, but in fact no psychometric properties of the SARS were provided (e.g., sample size, sensitivity, and

specificity of the test). As a result, the diagnosis of developmental dyslexia in Russian-speaking children based on this test remains a challenge to clinical practice.

We conducted a study aiming to overcome some current limitations of the SARS and to improve its clinical application. First, to verify the relevance of the existing normative data we collected new data on the SARS performance by typically developing Russian school-children with measured intact non-verbal intelligence and normal hearing, normal or corrected to normal vision, and no diagnosed neurological disorders. Unlike the authors of the SARS, who provided norms for grades 2 to 6 of the Russian school system, we also tested children in the second half of the grade 1, to make it possible to identify children with or at risk of developing reading impairment as early as possible. As a result, we report new direct data (means and standard deviations) for reading fluency and reading comprehension in typically developing Russian school-children for each of the primary school grades.

Additionally, we assessed the validity of the SARS for the diagnosis of developmental dyslexia. For that, we tested 50 children who were clinically diagnosed with developmental dyslexia. We examined the correspondence between the clinical diagnosis and the performance of these children on the SARS using original norms provided in the last published manual for the test (Kornev & Ishimova, 2010), and our new data.

Method

Participants

Typically developing participants were recruited at three Russian public schools, two in Moscow ($n=58$) and one in Volgograd ($n=47$). All children (total $N=105$) were native Russian speakers, 7 to 11 years of age, and were elementary school students in the first through fourth grades. The children had no history of diagnosed neurological disorders, no diagnosed problems with reading acquisition, and all had normal or corrected-to-normal vision (as reported in the informed consents given by their parents or legal representatives). Screening for primary auditory impairments (using the Audiogramm program version 4.6.1.3, Professional Audiometric System; Sennheiser HDA 280 audiometry headphones) resulted in exclusion of three participants. Screening for non-verbal intelligence with the Raven's Colored Progressive Matrices (Raven, 2004; Raven, Raven, & Kort, 2012) resulted in the exclusion of 12 participants who scored below norms. Ninety children (48 girls; 50 children from Moscow; 7 left-handers; $M_{age}=8.7$, $SD=1.13$) remained in the analysis. General information about typically developing participants included in the analysis is presented in Table 1.

Before using the test results of children from different cities in the consolidated analysis, we analyzed the effect of the city on reading performance, taking into account the grade and gender of participants. For this, we used linear regression models that were built with the *lme4* package for the statistical data processing program *R* (Bates et al., 2015). Factors such as city, grade, and gender were included in the model as predictors for reading fluency. No statis-

Table 1. Overview of Typically Developing Participants

N=90	First graders		Second graders		Third graders		Fourth graders	
	girls	boys	girls	boys	girls	boys	girls	boys
Total (M/V)	12 (12/0)	6 (6/0)	15 (11/4)	12 (7/5)	13 (7/6)	13 (5/8)	8 (1/7)	11 (1/10)

Note. (M/V) — the number of participants tested in Moscow / in Volgograd.

tically significant effect of the *city* factor on reading speed was found ($Est=6.681$, $SE=3.94$, $t=1.69$, $Pr(>|t|)=.094$). In a similar model, factors such as city, grade and gender were also included as predictors for the number of correctly answered questions. Again, we found no statistically significant effect of the *city* factor on the level of reading comprehension ($Est=-0.873$, $SE=0.575$, $t=-1.52$, $Pr(>|t|)=.133$).

All participants with developmental dyslexia ($N=50$, girls = 17; 1 left-hander; $M_{age}=8.9$, $SD=1.2$) were native Russian speakers. They were elementary school students, first through fourth graders, and, based on clinical assessment, had oral language skills typical for their age. The inclusion criterion was the confirmation of dyslexia by clinical specialists of the Center for Speech Pathology and Neurorehabilitation (Moscow) right before the study. This involved a certified speech therapist and neuropsychologist of the children's department assessing the children's speech development and other higher mental functions. It is worth noting that only children whose parents applied to the Center of Speech Pathology and Neurorehabilitation on their own initiative were invited to participate in the study. These parents reported persistent difficulties in their children with acquiring reading, which they tried to resolve either on their own or with the help of school or private speech therapists (this parameter varied), but did not achieve a persistent sufficient effect.

According to the neuropsychological assessment (Akhutina et al., 2016), the children with developmental dyslexia were heterogeneous in terms of the types of deficits observed. Some of the children had difficulties with processing visual and visual-spatial information, some children showed decreased skills in phonological processing, verbal memory, or the perception of rhythms, and some of the children displayed rapid exhaustion from activities related to the perception of written text. In some cases, a combination of two or more deficits with varying severity was observed. The quality of the children's writing ranged from nearly normal (when rewriting the text from the sample) to almost completely absent. Additional inclusion criteria were normal or corrected-to-normal vision, normal hearing, normal non-verbal intelligence (measured with the same instruments as in the typically developing children), and the absence of the history of neurological disorders.

Written informed consent forms were signed by parents or legal representatives of the children; children also orally agreed to participate. The study was approved by the Committee on Interuniversity Surveys and Ethical Assessment of Empirical Research, National Research

University Higher School of Economics, Russia (protocol No. 47 at 08.05.2018).

Materials and Procedure

Children's reading fluency and reading comprehension were assessed with the SARS (Kornev & Ishimova, 2010). For the initial testing, the authors recommended using the first halves of Texts I and II (three paragraphs from each text). "In secondary assessments it is possible to offer the child to read the second halves of texts I or II or to use the additional texts given in Appendix 2" (Kornev & Ishimova, 2010). Each half of the text has its own 10 questions for assessing reading comprehension. We used the first halves of Texts I and II ("How I caught crayfish" and "Ungrateful spruce") and their questions.

The first text (3 paragraphs, 97 words) was simpler; the second text (3 paragraphs, 127 words) was more difficult according to the authors of the methodology (Kornev & Ishimova, 2010). To quantify the level of difficulty of Text I and Text II, we analyzed the words included in terms of frequency (in instances per million) and length (in syllables and in letters). We used information on the frequency of words from the linguistic database for Russian words "StimulStat Project" (Alexeeva, Slioussar, & Chernova, 2017). When counting, we took into account all full-meaning words (by lemmas). Words that occur in the text several times were counted by the number of occurrences. The results of the analysis confirmed that Text I contains higher frequency words, while Text II uses longer words (see Table 2).

Table 2. Properties of Text I and Text II

	Text I	Text II
Average frequency of full-meaning lexical words (ipm)	2249.5	1102.0
Average length of full-meaning words in syllables	1.94	2.15
Average length of full-meaning words in letters	4.90	5.88

All children were informed in advance that after reading each text, they would have to answer questions to the text. Among our tested cohort of typically developing children there were no first graders who could not read the proposed texts. Considering that the SARS is the only standardized method for assessing reading skills in Russian, and that it is widely used by speech therapists, we cannot strictly claim that all children were seeing these texts for the first time. However, none of the children involved in the testing reported after reading that she or he had already read the text. These texts were not used to test reading skills when selecting children with dyslexia for our study.

According to the original procedure, children had to read the texts aloud and to answer ten content questions immediately after reading each of the texts. The vast majority of children could not read more than three paragraphs in one minute. If the child read slowly, we did not stop him until he had finished reading the third paragraph — that is, a fragment sufficient to answer all the questions. If the

child read quickly, we did not interrupt him at the end of reading the third paragraph, and waited until one minute had passed from the moment when he finished reading the title of the text, and only then stopped, to be able to count the number of words correctly read in one minute.

According to original guidelines, the examiner should listen to a child's reading and control the reading parameters using the form with the number of words printed at the end of each line. However, for some children it was distractive, and we audio-recorded children while they were reading. These records were further analyzed off-line.

Reading fluency was measured as the number of words read accurately in the first minute (the total number of words read minus the number of incorrectly read words (Kornev & Ishimova, 2010)). Reading comprehension was measured as the number of correct responses to the questions. When a participant could not read the second text at all (10 children with dyslexia), we considered reading fluency equal to 0, the level of reading comprehension equal to 0, and the number of incorrectly read words was marked as NA (no data).

For diagnostic purposes, the authors (Kornev, 1997, 2003; Kornev & Ishimova, 2010) suggested to transform raw reading fluency scores to a coefficient of reading technique (CoRT) by the following formula:

$$\text{CoRT} = 100 + ((M - m) / m) * 50,$$

where M is the number of words read correctly by a child in the first minute, m is the mean number of words read correctly in one minute by children of the same grade. The correspondence between reading fluency and CoRTs for particular grade was presented in the CoRTs table (Kornev & Ishimova, 2010). As an argument for converting direct data on children's reading fluency into CoRTs, the authors of the SARS mentioned the possibility of bringing the data to a standard scale that coincides in dimension with the IQ scale used in ABM-WISC, the adapted version of Wexler's test (Panasyuk, 1973). This was to provide the convenience of comparing the results of children tested with SARS and ABM-WISC. However, at present, Wexler's test for Russian-speaking children has limitations as a method of psychometry (Bazylchik, 2016). The comparison of reading test results with intelligence test results is extremely important for identifying children with dyslexia, but it is necessary to use methods with actual normative data. Considering the fact that normative data can be outdated, and that the restandardization of tests can take place at different rates, in our opinion, it is expedient to provide normative data on reading that indicate sample means and standard deviations, so that researchers and practitioners have the opportunity to choose the best intelligence tests available at a particular time. In our study we used the test of J. Raven, because it has more relevant norms for Russian-speaking children, which are given in the official guidelines (Raven et al., 2012) and which were collected with the participation of Russian-speaking primary schoolchildren (Davydov & Chmyhova, 2016).

The authors of the SARS recommended considering the performance within one standard deviation lower from the mean reading fluency for a particular grade as non-pathological forms of reading delay, which can be

presumably corrected by pedagogical efforts. In turn, reading fluency lower than one and a half standard deviations from the mean reading fluency was supposed to be considered as belonging to dyslexia spectrum. The values in between one and one and a half standard deviations would signal the risk of dyslexia. We used the above-mentioned recommendations to assess the validity of the SARS (Kornev & Ishimova, 2010).

Due to the lack of other standardized tests for reading in Russian, it was not possible to assess that type of *criterion validity*, which is calculated as the correlation coefficient between the scores on the test under study and an external measure — results of similar tests acknowledged as the “gold standard” (Ivanova & Hallowell, 2013). Therefore, we used a type of criterion validity known as *concurrent validity*, which is calculated by comparing it with another criterion — the presence or absence of a clinical diagnosis of developmental dyslexia in the child participant. We analyzed the data of 90 typically developing children and 50 children clinically diagnosed as having developmental dyslexia. We assessed such psychometric properties of the SARS as sensitivity (the percentage of children with clinically diagnosed developmental dyslexia who perform below a cutoff score for normal performance and who would be diagnosed as having dyslexia while using SARS), and specificity (the proportion of participants without developmental dyslexia who obtain results above the cutoff for normal reading abilities and who would be considered as typically developing based on the SARS; Ivanova & Hallowell, 2013). Additionally, we examined the correspondence between the clinical diagnosis and the performance of these children on the SARS using our new data for typically developing children.

Results

Reading Fluency

Table 3 shows the mean reading fluency for the typically developing children from grades 1 to 4. For comparison, we present the data published by the author of the SARS in 1997 (the data re-published in 2003 was the same as in 1997). The manual published in 2010 did not contain means and standard deviations for reading fluency, but reported the CoRT formula transformations of raw norms (see Method section). Therefore, we took the formula of CoRT from Kornev (2003), and based on the CoRT table in Kornev and Ishimova (2010), we calculated the average reading fluency for children from grades 2 to 4, which should have been used to get such CoRTs.

The authors of the SARS manual (Kornev & Ishimova, 2010) suggested that non-pathological forms of reading delay should be considered within one standard deviation lower from the mean reading fluency of the correspondent grade, and as belonging to dyslexia spectrum should be considered results lower than one and a half standard deviations from the mean (see Method section). We used the same cut-offs for our new data. It was not possible to reliably reconstruct standard deviations for the Kornev and Ishimova sample (2010), since the authors did not provide raw individual data in open sources. Therefore, we present

Table 3. Mean Reading Fluency

Grade	Mean Reading Fluency (Kornev, 1997, 2003)		Calculated Means from the CoRTs (Kornev & Ishimova, 2010)		Mean Reading Fluency from Our New Data	
	Text I	Text II	Text I	Text II	Text I	Text II
1	NA	NA	NA	NA	43.3	31.2
2	52.2	41.7	62.5	52.0	70.3	55.1
3	101.4	89.5	84.7	72.5	90.4	75
4	96.1	98.5	106.4	92.6	94.5	95.4

Note. NA — data were not provided.

the comparison of mean reading fluency and standard deviations between our normative data and the data published by Kornev (1997, the same in 2003) in Table 4. In Table 5, we present the reference levels for assessing reading fluency based on our data: normative, at risk, and dyslexia values.

Table 4. Comparison of Mean Reading Fluency and Standard Deviations of Reading Fluency Provided in Kornev (1997, 2003) and from Our Data

Grade	Kornev Data				Our New Data			
	Text I		Text II		Text I		Text II	
	Mean Reading Fluency	SD	Mean Reading Fluency	SD	Mean Reading Fluency	SD	Mean Reading Fluency	SD
1	NA	NA	NA	NA	43.3	13.5	31.08	10.1
2	52.2	11.2	41.7	14.9	70.3	16.6	55.1	15.1
3	101.4	29.3	89.5	23.7	90.4	15.8	75.0	17.1
4	96.1	27.7	98.5	27.8	94.5	7.8	95.4	15.2

Note. SD — standard deviation.

Reading Comprehension

The second important criterion for assessing reading skills is the level of reading comprehension. For this purpose, the author of the SARS offered 10 content questions to each of the texts. It was recommended to evaluate the results at 4 levels: A — 10 out of 10 correct answers, B — 7 to 9 correct answers, C — 4 to 6 correct answers, and D — 3 or less correct answers. Levels A and B should be considered

Table 5. Reference Levels for Assessing Reading Fluency (Based on Our New Data)

Typically Developing Children, N=90						
Grade	Text I			Text II		
	Mean (typical reading)	Mean -1 SD (at risk)	Mean -1.5 SD (dyslexia)	Mean (typical reading)	Mean -1 SD (at risk)	Mean -1.5 SD (dyslexia)
1	43.3	29.8	23.2	31.8	21.7	16.6
2	70.3	53.7	45.4	55.1	40.0	32.4
3	90.4	74.6	66.6	75.0	57.9	49.2
4	94.5	86.7	82.8	95.4	80.2	72.5

Note. *SD* — standard deviation.

as “functionally sufficient”, that is allowing to use the full information contained in a text. However, mean data for the number of correctly answered questions to the texts in the typically developing group were never published by Kornev and Ishimova, thus we cannot compare the results of reading comprehension based on our data with their findings.

In Table 6, we present reference levels for assessing reading comprehension based on the performance of 90 TD children in our tested cohort: normative, at risk, and dyslexia values. These data show that even in the fourth grade, typical children with normal non-verbal intelligence give an average of about 5 correct responses to 10 questions for Text II (the younger ones give even fewer correct responses). Therefore, the original criterion recommended by the SARS manual (Kornev & Ishimova, 2010), that is, 7 to 10 correct answers, can hardly serve as a basis for diagnosing reading comprehension impairment.

Table 6. Reference Levels for Assessing Reading Comprehension (Based on Our New Data)

Typically Developing Children, N=90						
Grade	Text I			Text II		
	Mean (typical reading)	Mean -1 SD (at risk)	Mean -1.5 SD (dyslexia)	Mean (typical reading)	Mean -1 SD (at risk)	Mean -1.5 SD (dyslexia)
1	6.17	3.9	2.7	3.6	1.3	0.2
2	7.0	5.5	4.8	4.3	2.2	1.1
3	7.8	6.3	5.6	4.9	2.7	1.6
4	7.3	5.9	5.3	5.3	3.4	2.5

Note. *SD* — standard deviation.

Validity of the SARS

As described in the Method section, we examined the validity of the SARS (Kornev & Ishimova, 2010) for the diagnosis of developmental dyslexia, calculating psychometric properties of the test such as sensitivity and specificity using original norms provided in the last published manual for the test (Kornev & Ishimova, 2010), and our new data for typically developing children.

Specificity

To assess the specificity of the test, we analyzed what percentage of typically developing children would fall into the typically developing group based on the SARS (Kornev & Ishimova, 2010). Since the normative data on the SARS were published only for the second and subsequent grades, we could not evaluate the results of 18 first-graders from our sample. Among 72 children in our group of typically developing schoolchildren in grades 2 to 4, there was not a single child whose CoRT would fall into the dyslexia group (that is, below 1.5 SDs), according to the CoRT table published in Kornev and Ishimova (2010). In 69 children, the CoRT values were within 1 SD (typical reading), and in three children (one fourth grader and two second graders) — between 1 and 1.5 SDs (risk group). According to these results, no child would be erroneously diagnosed as having dyslexia, and specificity of the SARS can be assessed as 100%.

It is worth noting that in the absence of screening for non-verbal intelligence, vision and hearing, non-specific reading disorders can be erroneously attributed to dyslexia, while the problem may be caused by another comorbid impairment (Snowling, Nash, Gooch, Hayiou-Thomas, & Hulme, 2019). Our initial cohort contained 15 children (further excluded from the analysis) with no dyslexia diagnosis, but with registered hearing impairments or performing below norms on the non-verbal intelligence test (Raven, 2004). In case we did not screen them for these exclusion criteria, four children (two children for both texts and two other children for Text II) could be erroneously attributed to the dyslexia group based on their CoRT values, which would reduce the specificity of the test.

We checked how the specificity of the test would change when using diagnostic criteria based on our new data. Out of 90 students of grades 1 to 4 from our group of typically developing schoolchildren, the reading performance of five children (two second graders, one child from the third grade and two children from the fourth grade) were below the level of 1.5 standard deviations from the average. That is, these children would fall into the dyslexia group, which would reduce the specificity of the test from 100% to 94.4%.

Sensitivity

To assess the sensitivity of the test, we analyzed what percentage of children with clinically diagnosed developmental dyslexia fell into the dyslexia group based on the SARS (Kornev & Ishimova, 2010). Nine children of the first grade could not be assessed, since the normative data for the SARS were only published starting from the second grade. Table 7 shows the percentage distribution of the remaining 41 children of grades 2 to 4 among typically reading, at risk and dyslexia groups, based on the CoRT for Text I and Text II.

Table 7. Distribution of Children with Clinically Diagnosed Dyslexia into Groups Based on the SARS and Norms from Kornev and Ishimova (2010) (Grades 2 to 4)

N=41	Typical Reading Fluency, %	At Risk, %	Dyslexia, %
Text I	51.2	12.2	36.6
Text II	43.9	17.1	39.0

Thus, according to the reading performance in Text I, 15 out of 41 children (from grades 2 to 4) with clinically diagnosed dyslexia would be classified as having dyslexia using normative data published in 2010. This gives a test sensitivity of 36.6%. According to the reading performance in Text II, 16 children would be classified as having dyslexia, which corresponds to a test sensitivity of 39.0%.

We investigated how the situation would have changed if we used the cutoff levels according to our new data obtained in 2018 (summarized in Table 5). Table 8 shows the percentage distribution of all tested children with clinically diagnosed dyslexia into groups based on the SARS performance in reading test and according to the new data for Text I and Text II.

Table 8. Distribution of Children with Clinically Diagnosed Dyslexia into Groups Based on the SARS and Our New Data for Reading Fluency (Grades 1 to 4)

N=50	Typical reading fluency, %	At risk, %	Dyslexia, %
Text I	28.0	0.0	72.0
Text II	26.0	14.0	60.0

With the new cutoff values for reading Text I, 34 out of 50 children with clinically diagnosed dyslexia would be classified as having dyslexia using our normative data. This gives a test sensitivity of 72.0%. According to the reading fluency of Text II, 30 out of 50 children with a diagnosis of developmental dyslexia would be classified as dyslexic, which corresponds to a sensitivity of 60.0%.

We also examined if the sensitivity value of the SARS would have changed if, in addition to the reading performance, reading comprehension would have been assessed. We used the same diagnostic cutoff values as were suggested by original guidelines for reading fluency: one standard deviation from the mean for the at-risk performance, and one and a half standard deviations for dyslexic performance. We used only our new data (see Table 6), because the authors of the SARS have not published normative data for reading comprehension. In our tested cohort, all participants with impaired reading comprehension had impaired reading fluency, but there were 13 children who had typical comprehension combined with impaired reading fluency. Table 9 shows the percentage distribution of all tested children with clinically diagnosed dyslexia among groups based on the SARS performance in

reading comprehension and according to our data for Text I and Text II.

Table 9. Distribution of Children with Clinically Diagnosed Dyslexia Among Groups Based on the SARS and Our New Data for Reading Comprehension (Grades 1 to 4)

N=50	Typical reading comprehension, %	At Risk, %	Dyslexia, %
Text I	78.0	2.0	20.0
Text II	62.0	0.0	38.0

Among 50 children with clinically diagnosed dyslexia, 10 performed below 1.5 *SD* answering the questions to Text I. These children would be diagnosed as dyslexic if only reading comprehension performance were used as diagnostic criterion. This gives a test sensitivity of 20%. 19 children performed below that cutoff value answering the questions to Text II, which corresponds to a test sensitivity of 38%.

We performed an additional analysis in search of an explanation for why 14 children (28% of all dyslexic children tested), who were within the normative range in terms of reading fluency and reading comprehension performance based on the SARS, were clinically diagnosed as dyslexic. For that, we listened again to the audio recordings of children's reading and weighted types of errors using a system of penalty points. Kornev and Ishimova (2010) claimed that

Table 10. System of Penalty Points for Different Types of Errors

Types of Errors	Penalty Points	
	Without Self-Correction	With Self-Correction
Wrong stress in a word	1	0.5
One sound skipped, changed or added	1	0.5
2–3 sounds swapped	2	1
2–3 sounds skipped, changed or added	2	1
One word skipped	2	1
One word repeated	2	1
One functional word added	2	1
One content word added	3	1.5
One word changed (4 or more sounds changed in a word)	3	1.5
Lost line with disorientation	3	1.5

“the errors that made children with dyslexia do not qualitatively differ from the errors that are normally observed in beginner readers.” In our tested cohort children made errors at different levels: from minor, such as repeating one sound or reducing sounds in a weak position, to major, such as saying another word instead of the one presented, or losing a line and disorienting. Table 10 presents the types of errors that occurred in the typically developing and dyslexic groups, along with the system of penalty points for these errors.

For each child, we calculated the total weighted score for errors by summing the penalty points for each error made. The results of 90 typically developing children are shown in Table 11. Then, we calculated the penalty points for each of those 14 children with a clinical diagnosis of dyslexia but no reading deficit identified by the main scores of the SARS. We found that for all of these children, the penalty points were more than 1.5 standard deviation higher than the means for the corresponding grade for at least one of the texts. In other words, these children read fluently but with a lot of errors of greater weight. Thus, they have an impairment of reading quality, not speed, which remains unnoticed when using the original scoring criteria for the SARS (Kornev & Ishimova, 2010).

Table 11. Reference Levels for Evaluation the Weighted Error Scores (Based on Our New Data)

Typically Developing Children, <i>N</i> =90						
Text I			Text II			
Grade	Mean Error Score	Mean + 1 SD	Mean + 1.5 SD	Mean Error score	Mean + 1 SD	Mean + 1.5 SD
1	2.75	4.61	5.54	2.72	4.70	5.69
2	3.77	7.15	8.84	4.44	7.69	9.32
3	3.40	5.92	7.18	5.54	9.56	11.57
4	2.95	5.90	7.38	5.76	9.52	11.4

Note. *SD* — standard deviation.

We explored whether the sensitivity of the SARS would change if, in addition to reading fluency, reading quality was assessed taking into account a new criterion — weighted penalty points for errors. We used the same diagnostic cutoff levels that were proposed in the original manual for reading speed: one standard deviation from the mean for the results considered as a risk group, and one and a half standard deviations from the mean for the results related to dyslexia. We used our data for 50 children with dyslexia for Text I, but for 40 children with dyslexia for Text II, since 10 children with dyslexia (20% of participants with dyslexia) could not read Text II at all, and there are no data about errors in the second text for them. Table 12 shows the percentage distribution of all tested children with clinically diagnosed dyslexia into groups based on weighted error scores in the SARS in accordance with our data for Texts I and II.

Table 12. Distribution of Children with Clinically Diagnosed Dyslexia into Groups Based on Weighted Error Scores in SARS (in Accordance with Our Data, Grades 1 to 4)

	Weighted Error Scores Within Normative Range, %	At Risk, %	Dyslexia, %
Text I (<i>N</i> =50)	30.0	6.0	64.0
Text II (<i>N</i> =40)	7.5	15.0	77.5

Among 50 children with clinically diagnosed dyslexia, 32 showed results below the cutoff level of 1.5 standard deviations when reading Text I. These children would be diagnosed as having developmental dyslexia if only weighted errors scores were used as a diagnostic criterion. This gives a test sensitivity of 64%. Among the 40 children with dyslexia who could read the second text, 31 children showed results below the cutoff level, which corresponds to a test sensitivity of 77.5%.

Discussion

Standardized tests should be used with actual normative data for diagnostic purposes, because normative performance can change over time (Raven, 2000). The fact that normative data need to be updated concerns not only this particular methodology but is characteristic of most standardized tests. In the methodical manual published in 2010 (Kornev & Ishimova, 2010), the authors also noted the reasonability of normative data updating. Importantly, for appropriate cutoff scores for normal performance it is preferable to test not randomly chosen children, but those who match such inclusion criteria as normal intellectual abilities, vision and hearing, and the absence of diagnosed neurological disorder (Ivanova & Hallowell, 2013).

The data we collected for the SARS in 2018 provide an update of reference levels for assessing reading performance and are controlled for the mentioned inclusion criteria for the tested normative sample. Although the number of typically developing children participating in our study (*N*=90) is not enough for complete restandardization of the test, the fact that the norms calculated on a sample of 90 children were diagnostically more productive than the norms from the author’s manual (Kornev & Ishimova, 2010) to identify a clinical phenomenon, suggests that it is necessary to review the normative levels.

As for the diagnostic power of the SARS, the specificity of the test is perfect. That means that any normative data set of SARS allows the identification of typical reading performance and the absence of dyslexia in a new cohort with high accuracy: from 100% (when using normative data obtained in 2007–2008) to 94.4% (when using reference levels, calculated on the data of 2018). However, sensitivity of the SARS substantially depends on the normative data used for cutoff values. When normative data obtained in 2007–2008 were used, sensitivity of the test lied between 36.6% and 39.0% (depending on the text), while using our new normative data obtained in 2018, the sensitivity of the test reached 60.0–72.0%. A test sensitivity of even 72.0% means that 28.0% of children with impaired reading skills

may be erroneously not diagnosed as such. As it was shown in many studies, the consequence of the non-recognition of dyslexia is the lack of adequate intervention, leading to anxiety and depressive behaviors (Mugnaini, Lassi, La Malfa, & Albertini, 2009; Willcutt & Pennington, 2000; Nelson & Harwood, 2011; Törő, Miklósi, Horanyi, Kovács, & Balázs, 2018), suicidal ideation, school failure, and drop out (Barbiero et al., 2019; Daniel et al., 2006; Wilson, Armstrong, Furrie, & Walcot, 2009).

The relatively low sensitivity of the SARS is a result of using only one (albeit the most important) diagnostic criterion — the speed of decoding. However, the tool developed by Kornev and Ishimova in fact allows us to use two additional diagnostic criteria: the level of reading comprehension, and weighted errors scores. To further improve the clinical validity of the SARS, we provide reference levels for assessing reading comprehension and weighted error scores based on the results of our tested cohort of typically developing children.

Our study showed that the insertion of a weighted error assessment as an additional diagnostic criterion can improve the diagnostic validity of the SARS. We were able to calculate the penalty points for errors due to the changes in the testing procedure (we audio recorded children's voice while reading). Considering that this innovation solves the problem of children's distracted attention and allows for more accurate assessment of reading parameters, we can recommend the inclusion of an audio recording of the children's voices while reading in the diagnostic procedure.

As for the assessment of reading comprehension (which is carried out in the SARS by means of content questions which children answer after reading the texts), the results of these tests are rather variable, but on the whole are quite low even in the typically developing group. This leads to the fact that only a very low level of reading comprehension is diagnosed as impairment (see Table 6, as well as comments on page 16 in the author's manual of Kornev and Ishimova, 2010). Actual data show that the application of a reading comprehension criterion does not lead to an increase in the overall sensitivity of the test. This suggests either that this test can be useful only in specific cases, when a child has a very low level of reading comprehension combined with normal reading fluency and accuracy (in our sample we had no such children), or that the diagnostic power of this task is doubtful. However, since reading comprehension is extremely important for schooling in general, future research might investigate variants of more effective reading comprehension tests.

It is important that three aforementioned deficits (slow reading, lack of understanding of what is read, and numerous major errors while reading) measured with the SARS' metrics (reading fluency, reading comprehension and weighted error scores, correspondingly) may occur in children with dyslexia both in isolation and in combination. Our study showed that children with only one type of deficit — those with high speed of reading and normal reading comprehension, but with a large number of errors — were still clinically diagnosed as having developmental dyslexia.

We propose the use of not a single criterion but two or better yet three criteria for diagnostic purposes — reading speed, reading comprehension, and weighted error scores —

since none of the criteria separately provides test sensitivity close to 100 % (see the Results section). Reference levels are presented in Table 5, Table 6 and Table 11, correspondingly. If a child's performance is one and a half standard deviation lower than the mean for the same grade, developmental dyslexia should be diagnosed. Additionally, screening for non-verbal intellectual abilities, hearing and vision should necessarily accompany reading abilities testing, since these measures interact with the assessment of reading. It is worth noting that western researchers also use tests for reading words (see for example Brus & Voeten, 1973) and for reading pseudowords (see for example van den Bos, Spelberg, Scheepstra, & DeVries, 1994). The development of such standardized instruments for the Russian language and investigation of their effectiveness in the diagnosis of dyslexia is one of the possible directions for future studies.

In conclusion, we express our great respect to A.N. Kornev and O.A. Ishimova for developing the only available standardized test for the assessment of reading skills in Russian-speaking children. We hope that our new data, their clinical evaluation, and suggested amendments for scoring the results (concerning a more precise evaluation of reading comprehension and the introduction of weighted error scores) will expand the efforts of the authors of the methodology and will contribute to an even wider use of the SARS.

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